

University of Groningen

## Exposure-Response Analyses of Asbestos and Lung Cancer Subtypes in a Pooled Analysis of Case-Control Studies

Olsson, Ann C; Vermeulen, Roel; Schüz, Joachim; Kromhout, Hans; Pesch, Beate; Peters, Susan; Behrens, Thomas; Portengen, Lützen; Mirabelli, Dario; Gustavsson, Per

*Published in:*  
Epidemiology

*DOI:*  
[10.1097/EDE.0000000000000604](https://doi.org/10.1097/EDE.0000000000000604)

**IMPORTANT NOTE:** You are advised to consult the publisher's version (publisher's PDF) if you wish to cite from it. Please check the document version below.

*Document Version*  
Publisher's PDF, also known as Version of record

*Publication date:*  
2017

[Link to publication in University of Groningen/UMCG research database](#)

### *Citation for published version (APA):*

Olsson, A. C., Vermeulen, R., Schüz, J., Kromhout, H., Pesch, B., Peters, S., Behrens, T., Portengen, L., Mirabelli, D., Gustavsson, P., Kendzia, B., Almansa, J., Luzon, V., Vlaanderen, J., Stücker, I., Guida, F., Consonni, D., Caporaso, N., Landi, M. T., ... Straif, K. (2017). Exposure-Response Analyses of Asbestos and Lung Cancer Subtypes in a Pooled Analysis of Case-Control Studies. *Epidemiology*, 28(2), 288-299. <https://doi.org/10.1097/EDE.0000000000000604>

### **Copyright**

Other than for strictly personal use, it is not permitted to download or to forward/distribute the text or part of it without the consent of the author(s) and/or copyright holder(s), unless the work is under an open content license (like Creative Commons).

The publication may also be distributed here under the terms of Article 25fa of the Dutch Copyright Act, indicated by the "Taverne" license. More information can be found on the University of Groningen website: <https://www.rug.nl/library/open-access/self-archiving-pure/taverne-amendment>.

### **Take-down policy**

If you believe that this document breaches copyright please contact us providing details, and we will remove access to the work immediately and investigate your claim.

Downloaded from the University of Groningen/UMCG research database (Pure): <http://www.rug.nl/research/portal>. For technical reasons the number of authors shown on this cover page is limited to 10 maximum.

# Exposure–Response Analyses of Asbestos and Lung Cancer Subtypes in a Pooled Analysis of Case–Control Studies

Ann C. Olsson,<sup>a,b</sup> Roel Vermeulen,<sup>c</sup> Joachim Schüz,<sup>a</sup> Hans Kromhout,<sup>c</sup> Beate Pesch,<sup>d</sup> Susan Peters,<sup>c,e</sup> Thomas Behrens,<sup>d</sup> Lützen Portengen,<sup>c</sup> Dario Mirabelli,<sup>f</sup> Per Gustavsson,<sup>b</sup> Benjamin Kendzia,<sup>d</sup> Josue Almansa,<sup>c</sup> Veronique Luzon,<sup>a</sup> Jelle Vlaanderen,<sup>a,c</sup> Isabelle Stücker,<sup>g,h</sup> Florence Guida,<sup>g,h</sup> Dario Consonni,<sup>i</sup> Neil Caporaso,<sup>j</sup> Maria Teresa Landi,<sup>j</sup> John Field,<sup>k</sup> Irene Brüske,<sup>l</sup> Heinz-Erich Wichmann,<sup>l</sup> Jack Siemiatycki,<sup>m</sup> Marie-Elise Parent,<sup>n</sup> Lorenzo Richiardi,<sup>f</sup> Franco Merletti,<sup>f</sup> Karl-Heinz Jöckel,<sup>o</sup> Wolfgang Ahrens,<sup>p</sup> Hermann Pohlabein,<sup>p</sup> Nils Plato,<sup>b</sup> Adonina Tardón,<sup>q</sup> David Zaridze,<sup>r</sup> John McLaughlin,<sup>s</sup> Paul Demers,<sup>t</sup> Neonila Szeszenia-Dabrowska,<sup>u</sup> Jolanta Lissowska,<sup>v</sup> Peter Rudnai,<sup>w</sup> Eleonora Fabianova,<sup>x</sup> Rodica Stanescu Dumitru,<sup>y</sup> Vladimir Bencko,<sup>z</sup> Lenka Foretova,<sup>aa</sup> Vladimir Janout,<sup>bb</sup> Paolo Boffetta,<sup>cc</sup> Bas Bueno-de-Mesquita,<sup>dd</sup> Francesco Forastiere,<sup>ee</sup> Thomas Brüning,<sup>d</sup> and Kurt Straif<sup>a</sup>

Submitted 20 October 2015; accepted 22 November 2016.

From the <sup>a</sup>International Agency for Research on Cancer, Lyon, France; <sup>b</sup>The Institute of Environmental Medicine, Karolinska Institutet, Stockholm, Sweden; <sup>c</sup>Institute for Risk Assessment Sciences, Utrecht, The Netherlands; <sup>d</sup>Institute for Prevention and Occupational Medicine of the German Social Accident Insurance – Institute of the Ruhr-Universität Bochum (IPA), Bochum, Germany; <sup>e</sup>Occupational Respiratory Epidemiology, School of Population Health, University of Western Australia, Perth, Australia; <sup>f</sup>Cancer Epidemiology Unit, Department of Medical Sciences, University of Turin and CPO Piemonte, Turin, Italy; <sup>g</sup>INSERM, Centre for research in Epidemiology and Population Health (CESP), U1018, Environmental epidemiology of cancer Team, Villejuif, France; <sup>h</sup>Université Paris-Sud, UMRS 1018, Villejuif, France; <sup>i</sup>Epidemiology Unit, Fondazione IRCCS Ca' Granda—Ospedale Maggiore Policlinico, Milan, Italy; <sup>j</sup>National Cancer Institute, Bethesda, MD; <sup>k</sup>Roy Castle Lung Cancer Research Programme, The University of Liverpool Cancer Research Centre, Department of Molecular and Clinical Cancer Medicine, Institute of Translational Medicine, Liverpool, United Kingdom; <sup>l</sup>Institut für Epidemiologie, Deutsches Forschungszentrum für Gesundheit und Umwelt, Neuherberg, Germany; <sup>m</sup>University of Montreal Hospital Research Centre (CRCHUM), Montreal, Canada; <sup>n</sup>INRS-Institut Armand-Frappier, Université du Québec, Laval, Québec, Canada; <sup>o</sup>Institute for Medical Informatics, Biometry and Epidemiology, University of Duisburg-Essen, Essen, Germany; <sup>p</sup>Leibniz Institute for Prevention Research and Epidemiology – BIPS, Bremen, Germany; <sup>q</sup>The Biomedical Research Centre Network for Epidemiology and Public Health (CIBERESP), University of Oviedo, Oviedo, Spain; <sup>r</sup>Russian Cancer Research Centre, Moscow, Russia; <sup>s</sup>Public Health Ontario, Toronto, Canada; <sup>t</sup>Occupational Cancer Research Centre, Cancer Care Ontario, Toronto, Canada; <sup>u</sup>The Nofer Institute of Occupational Medicine, Lodz, Poland; <sup>v</sup>The M Skłodowska-Curie Cancer Center and Institute of Oncology, Warsaw, Poland; <sup>w</sup>National Centre for Public Health, Budapest, Hungary; <sup>x</sup>Regional Authority of Public Health, Banska Bystrica, Slovakia; <sup>y</sup>Institute of Public Health, Bucharest, Romania; <sup>z</sup>Institute of Hygiene and Epidemiology, 1st Faculty of Medicine, Charles University, Prague, Czech Republic; <sup>aa</sup>Masaryk Memorial Cancer Institute and Medical Faculty of Masaryk University, Department of Cancer Epidemiology & Genetics, Brno, Czech Republic; <sup>bb</sup>Faculty of Medicine, Palacky

University, Olomouc, Czech Republic; <sup>cc</sup>The Tisch Cancer Institute and Institute for Translational Epidemiology, Icahn School of Medicine at Mount Sinai, New York, NY; <sup>dd</sup>National Institute for Public Health and the Environment (RIVM), Bilthoven, The Netherlands; and <sup>ee</sup>Department of Epidemiology, ASL RomaE, Rome, Italy.

Joachim Schüz and Kurt Straif are part of a collaborative study with the State Russian Institute of Occupational Health (SRIOH) on asbestos and cancer risk, which includes financial support from SRIOH to IARC. Paolo Boffetta acted as expert witness for Edison SpA in asbestos-related litigation, outside the submitted work. Beate Pesch, Thomas Behrens, Benjamin Kendzia, and Thomas Brüning, as staff of the Institute for Prevention and Occupational Medicine (IPA), are employed at the “Berufsgenossenschaft Rohstoffe und chemische Industrie” (BG RCI), a public body, which is a member of the study’s main sponsor, the German Social Accident Insurance (DGUV). IPA is an independent research institute of the Ruhr-Universität Bochum. The authors are independent from the German Social Accident Insurance in study design, access to the collected data, responsibility for data analysis and interpretation, and the right to publish. The views expressed in this article are those of the authors and not necessarily those of the sponsor. The other authors report no conflicts of interest. The SYNERGY project is funded by the German Social Accident Insurance (DGUV), Grant FP 271. The original studies were funded as follows: in Canada by Canadian Institutes for Health Research and Guzzo-SRC Chair in Environment and Cancer, National Cancer Institute of Canada, Canadian Cancer Society, Occupational Cancer Research Centre, Workplace Safety and Insurance Board, Canadian Cancer Society, and Cancer Care Ontario; in France by the French Agency of Health Security (ANSES), Fondation de France, French National Research Agency (ANR), National Institute of Cancer (INCA), Fondation pour la Recherche Médicale, French Institute for Public Health Surveillance (InVS), Health Ministry (DGS), Organization for the Research on Cancer (ARC), and French Ministry of work, solidarity, and public function (DGT); in Germany by Federal Ministry of Education, Science, Research, and Technology (Grant 01 HK 173/0), Federal Ministry of Science (Grant 01 HK 546/8), and the Ministry of Labour and Social Affairs (Grant IIIb7-27/13); in Italy by Environmental Epidemiology Program of the Lombardy Region, INAIL, Italian Association for Cancer Research, Region Piedmont, Compagnia di San Paolo, and Lazio Region; in Poland by Polish State Committee for Scientific Research (Grant SPUB-M-COPERNICUS/P-05/ DZ-30/99/2000); in Czech Republic by MH CZ - DRO (MMCI, 00209805); in Spain by Instituto Universitario de Oncología, Universidad de Oviedo, Asturias, Fondo de Investigación Sanitaria (FIS) and Ciber de Epidemiología y Salud Pública (CIBERESP); in Sweden by Swedish Council for Work Life

Copyright © 2017 The Author(s). Published by Wolters Kluwer Health, Inc. This is an open access article distributed under the Creative Commons Attribution License 4.0 (CCBY), which permits unrestricted use, distribution, and reproduction in any medium, provided the original work is properly cited.

ISSN: 1044-3983/17/2802-0288

DOI: 10.1097/EDE.0000000000000604

Research and Swedish Environmental Protection Agency; in the Netherlands by Dutch Ministry of Health, Welfare and Sports, National Institute of Public Health and the Environment, and Europe Against Cancer Program; in the UK by Roy Castle Foundation; in the USA by Intramural Research Program of the National Institutes of Health, National Cancer Institute, Division of Cancer Epidemiology and Genetics, Bethesda, MD; the IARC multicenter study in Central and Eastern Europe was funded by the European Commission's INCO Copernicus program (Contract IC15-CT96-0313).

The data and the computer code are not available for replication because the data are not publicly available.

**SDC** Supplemental digital content is available through direct URL citations in the HTML and PDF versions of this article ([www.epidem.com](http://www.epidem.com)).

Correspondence: Ann C. Olsson, ENV/IARC, 150 cours Albert Thomas, 69372 Lyon Cedex 08, France. E-mail: [olssona@iarc.fr](mailto:olssona@iarc.fr).

**Background:** Evidence is limited regarding risk and the shape of the exposure–response curve at low asbestos exposure levels. We estimated the exposure–response for occupational asbestos exposure and assessed the joint effect of asbestos exposure and smoking by sex and lung cancer subtype in general population studies.

**Methods:** We pooled 14 case–control studies conducted in 1985–2010 in Europe and Canada, including 17,705 lung cancer cases and 21,813 controls with detailed information on tobacco habits and lifetime occupations. We developed a quantitative job–exposure–matrix to estimate job-, time period-, and region-specific exposure levels. Fiber-years (ff/ml-years) were calculated for each subject by linking the matrix with individual occupational histories. We fit unconditional logistic regression models to estimate odds ratios (ORs), 95% confidence intervals (CIs), and trends.

**Results:** The fully adjusted OR for ever-exposure to asbestos was 1.24 (95% CI, 1.18, 1.31) in men and 1.12 (95% CI, 0.95, 1.31) in women. In men, increasing lung cancer risk was observed with increasing exposure in all smoking categories and for all three major lung cancer subtypes. In women, lung cancer risk for all subtypes was increased in current smokers (ORs ~two-fold). The joint effect of asbestos exposure and smoking did not deviate from multiplicativity among men, and was more than additive among women.

**Conclusions:** Our results in men showed an excess risk of lung cancer and its subtypes at low cumulative exposure levels, with a steeper exposure–response slope in this exposure range than at higher, previously studied levels. (See video abstract at, <http://links.lww.com/EDE/B161>.)

(*Epidemiology* 2017;28: 288–299)

Asbestos is a general term for a group of mineral silicate fibers naturally found on all continents; the commercialized types are the serpentine mineral chrysotile (white asbestos) and the amphibole minerals amosite (brown asbestos), anthophyllite, crocidolite (blue asbestos), and tremolite.<sup>1</sup> Asbestos fibers are generally considered strong, flexible, stable, heat-resistant, and durable; they have therefore been attractive for a wide range of industrial applications for over a century. Consequently, large groups of workers have been (and still are, in many countries) exposed to asbestos, for example in the insulation, textile, cement, roofing, and refractory industries. The highest exposure levels have been measured

among workers manufacturing asbestos products or employed in asbestos mining and milling operations.<sup>2</sup> Asbestos has been banned successively since 1980s in many countries due to its adverse health effects.<sup>3</sup> Nevertheless, exposure may still occur when buildings insulated with asbestos are demolished, when asbestos is removed from any type of structure, and during maintenance and repair of asbestos-containing materials.<sup>2</sup> The World Health Organization (WHO) estimated in 2006 that 125 million workers worldwide are still exposed to asbestos.<sup>4</sup>

The International Agency for Research on Cancer (IARC) Monographs Programme evaluated the carcinogenicity of asbestos in 1973, 1977, 1987, and 2011; the Working Group concluded in the most recent evaluation (Vol. 100C) that all forms of asbestos cause mesothelioma and cancer of the lung, larynx, and ovary,<sup>3</sup> and made no distinction by lung cancer cell type when evaluating asbestos carcinogenicity to the lung.

Lung cancer is the most common cancer globally.<sup>5</sup> Tobacco smoking is well established as the main cause; for instance, in the United Kingdom, an estimated 85% of lung cancers in men and 47% of lung cancers in women are attributable to tobacco smoking.<sup>6</sup> Asbestos is the most important occupational carcinogen, and lung cancer is the most common asbestos-related cancer.<sup>7</sup>

Asbestos was the first occupational exposure to be suggested to have a joint effect with smoking.<sup>8</sup> Several studies and reviews have supported this hypothesis, but the type of interaction (additive or multiplicative) has been debated.<sup>9–13</sup>

Here, we used a pooled dataset of lung cancer case–control studies conducted in Europe and Canada (the SYNERGY project) to estimate lung cancer risk related to occupational asbestos exposure, and its interaction with smoking. The objectives of this work were to (1) estimate the lung cancer risk associated with quantitative indices of occupational asbestos exposure by sex, while adjusting for smoking; (2) assess the exposure–response relationship for asbestos and lung cancer by sex, major subtype, and smoking status; and (3) assess the joint effect of asbestos exposure and smoking on an additive and multiplicative scale.

## METHODS

### The SYNERGY Project

Fourteen case–control studies on lung cancer from Europe and Canada (see eTable 1 at <http://links.lww.com/EDE/B144>) were pooled in the SYNERGY project to study joint effects of occupational carcinogens, including asbestos, and smoking in relation to lung cancer risk. The studies LUCA and LUCAS were restricted to men, and PARIS to regular smokers with squamous-cell lung carcinoma (SQLC) and small-cell lung carcinoma (SCLC). Participation rates were 62%–98% (mean, 83%) among cases and 41%–100% (mean, 70%) among controls. All studies collected lifetime smoking histories and complete occupational histories, except

MORGEN. MORGEN is a case-control study nested in the European Prospective Investigation into Cancer and Nutrition (EPIC) study in the Netherlands, where 45% of those invited completed a questionnaire at recruitment.

The data were collected in 1985–2010, and almost all interviews with study participants were conducted face-to-face. LUCAS and MORGEN collected data using self-administered questionnaires, and women in MONTREAL and some participants in TORONTO were interviewed by phone. Next-of-kin were interviewed for most cases and some controls in LUCAS and some participants in ICARE and MONTREAL (9% of cases, 6% of controls). Controls were individually or frequency-matched to cases by sex and age, and mainly recruited from the general population (79%). Lung cancer subtypes were classified according to WHO guidelines after histological or cytological confirmation. Reference pathology was performed for the German cases.<sup>14</sup> Ethical approvals for the original studies were obtained in accordance with legislation in each country, and in addition from the IARC Ethics Committee. More information about the SYNERGY project is available at: <http://synergy.iarc.fr>.

Occupational data consisted of a list of employment periods for every study subject. For every period, job and industrial activity had been recorded, coded respectively to the International Standard Classification of Occupations from 1968 (ISCO-68) and the International Standard Classification of Industries, Revision 2, along with the start and end years.

### Assessment of Occupational Asbestos Exposure

Quantitative measurements of fibers (71,816) from 14 countries (mainly Germany, the UK, Canada, Italy, France, and Norway) were entered into the project-specific exposure database ExpoSYN according to a standardized protocol.<sup>15</sup> Most data points were determined by phase-contrast microscopy (>95%). It can be assumed that most data represented chrysotile (67%). Regarding measurement strategies, 53% of the measurements were considered “representative,” 9% “worst case,” and 38% “unknown.”<sup>15</sup> All measurements were linked to a standardized (ISCO-68) job title. Statistical models were applied to the personal measurements (27,958) collected in 1971–2009 to develop a project-specific quantitative job-exposure-matrix (SYN-JEM) for occupational asbestos exposure. Some measurements were attributed to jobs clearly unrelated to asbestos exposure, like teachers; we assumed these to represent exceptional situations, which should not be generalized to all individuals in that job. Therefore, a semi-quantitative general population job-exposure matrix based on ISCO-68 codes (DOM-JEM) was used in the model, where every job was rated as nonexposed (=0), low exposed with regard to exposure intensity or high exposed with low exposure probability (=1), or high exposed with high-exposure probability (=2). Jobs considered to be nonexposed in DOM-JEM were set to 0 fibers per milliliter (ff/ml) in SYN-JEM, disregarding actual measurements, if any. When there were <5

measurements for a specific job, the geometric mean estimate of all jobs within the same unit or major group was applied, so the job estimate was based on information from the most similar jobs. Because every job was expert-rated as being non-, low-, or high exposed, an exposure level for every potential job could be calculated, even in the absence of measurements for that particular job. Additional SYN-JEM model specifications and sensitivity analyses using alternative models are described elsewhere.<sup>16–18</sup> In brief, for all countries and occupations together, we implemented a linear historical trend with an annual decrease of fiber concentrations of –10.7% before ban implementation and no further downward trend after ban implementation, and an exposure ceiling before 1975 to avoid unrealistically high estimates due to unrestrained back-extrapolation to periods when actual measurements were not carried out. Linking the occupational histories of the participants to SYN-JEM generated individual job-, region-, and year-specific estimates of the average intensity of asbestos exposure during a standard 8-hour working day in ff/ml. Cumulative asbestos exposure (expressed as ff/ml-years) was defined as the average exposure intensity in a particular job multiplied by the years of employment, and totaled over the working life of the participants.

### Statistical Analyses

Unconditional logistic regression models were fitted to generate odds ratios (ORs) and 95% confidence intervals (CIs) of lung cancer associated with various indices of asbestos exposure. The subjects classified as nonexposed were the reference category in each of the analyses.

Three strategies for adjustment were applied: the first model (OR1) adjusted for age group (<45, 45–49, 50–54, 55–59, 60–64, 65–69, 70–74, 75+ years) and study; the second model (OR2) also adjusted for tobacco smoking as a continuous variable ( $\log[\text{cigarette pack-years} + 1]$ ) and for time-since-quitting smoking cigarettes (current smokers; quitting 2–7, 8–15, 16–25, 26+ years before diagnosis/interview; never-smokers); and the third model (OR3) also adjusted for ever-employment in a “list A” job (yes/no). “List A” is a list of occupations and industries known to present an excess risk of lung cancer, compiled by Ahrens and Merletti<sup>19</sup> and updated by Mirabelli et al.<sup>20</sup> Here, we modified “list A” so that jobs originally included solely because of asbestos exposure were excluded, to avoid potential over-adjustment. Examples are asbestos cement product makers, insulators, some jobs in mining and quarrying, and some jobs in manufacture of non-metallic mineral products not elsewhere classified (e.g., other crushers, grinders, and mixers; beam warpers; loom threaders; fabric examiners and repairers; spinners and winders).

Current smokers were people who had smoked >1 cigarette per day for >1 year, including those who had stopped smoking in the 2 years before diagnosis/interview. Cigarette pack-years were calculated as:  $\sum \text{duration} \times \text{average intensity per day}/20$ .



We used kernel plots to describe the distribution of cumulative asbestos exposure among cases and controls. Cumulative asbestos exposure in ff/ml-years was categorized according to quartiles of its distribution in controls for the main exposure–response analyses. In the analyses stratified by lung cancer subtype and smoking status, we used two exposure categories (below and above the median) because the number of observations was limited.

*P* values for linear trend were obtained by applying a logistic regression model including the respective continuous variable. The trend was calculated among all subjects and among exposed subjects only.

We examined robustness of results by sensitivity analyses as follows:

1. excluding one study at a time, to see if any specific study largely influenced the overall result;
2. excluding one industry at a time, to see if any specific industry largely influenced the overall result;
3. stratification by hospital- and population-based studies, to assess if associations differed by study design;
4. restricting the study base to blue-collar workers, to limit potential residual confounding from socioeconomic factors;
5. restricting the analyses to workers who started working in 1960 or later, as exposure data were scarce before the 1960s and exposure estimates in SYN-JEM may be affected by larger uncertainty;
6. excluding “laborers not elsewhere classified” (ISCO 9–99.10) because they represent a substantial proportion of exposed workers in some of the studies, to see if their inclusion had unduly influenced the results.

A multinomial logistic regression model and a likelihood ratio test were used to explore heterogeneity between the three major lung cancer subtypes in relation to a categorical variable of cumulative asbestos exposure.

Lagging of cumulative exposure was applied, in which exposure in the 5, 10, 15, or 20 years before diagnosis/interview was disregarded. As results did not differ by lag-times, we used unlagged models in the main analyses.

The slope of the exposure–response relationship reflects the average excess relative risk per fiber-year. It was obtained from a linear OR model adjusted for study, “list A” occupations, time-since-quitting smoking cigarettes, and smoking pack-year categories using maximum likelihood estimation and was expressed as  $K_L \times 100$ , that is, 100 times the excess relative risk per fiber-year.

We performed additional spline analyses using non-parametric smoothing as implemented in the R package *mgcv* to assess in more detail the shape of the exposure–response relationship. The optimal smoothing parameter was selected based on generalized cross-validation and under the assumption that the total degrees of freedom required for a biologically plausible model would not exceed 3. 95% CIs for ORs were derived by simulation from the posterior distribution

of the model coefficients, performing random draws from a multivariate normal distribution parameterized by the estimated mean vector and estimated covariance matrix of the model coefficients.

Meta-analyses were conducted to explore study-specific ORs using the Stata command “metan,” where the extent of heterogeneity between OR estimates was assessed as a percentage ( $I^2$ ).<sup>21</sup>

We assessed the additive interaction between smoking and asbestos by estimating the relative excess risk due to interaction.<sup>22</sup> We again used a linear OR model adjusted for covariates, and bootstrapped CIs for the excess risk due to interaction. Departure from multiplicative interaction between smoking and asbestos was assessed by testing the asbestos–smoking interaction term in the logistic model.

We conducted statistical analyses using SAS, version 9.3 (SAS Institute, Cary, NC); STATA, version 12.1 (Stata-Corp, College Station, TX); and R, version 3.2.

## RESULTS

We omitted study participants with incomplete data on covariates (804 cases, 848 controls), leaving 16,901 lung cancer cases (4,752 lung adenocarcinoma, 6,503 squamous cell carcinoma, 2,730 small-cell carcinoma, 2,822 other/unspecified lung cancer cell types, 94 not available) and 20,965 controls for the analyses.

### Characteristics of Subjects by Exposure Status

Table 1 shows characteristics of study participants by asbestos exposure status, disease status, and sex. Smoking status differed by asbestos exposure status; nonexposed were more often never-smokers among both men and women. Lung cancer pathology also differed by asbestos exposure status; adenocarcinoma was less frequent and squamous-cell carcinoma more frequent among asbestos-exposed compared with nonexposed men and women. More cases with asbestos exposure ever worked in other occupations with an anticipated lung cancer risk (20% in men) than controls (15% in men) or non-exposed cases (5% in men).

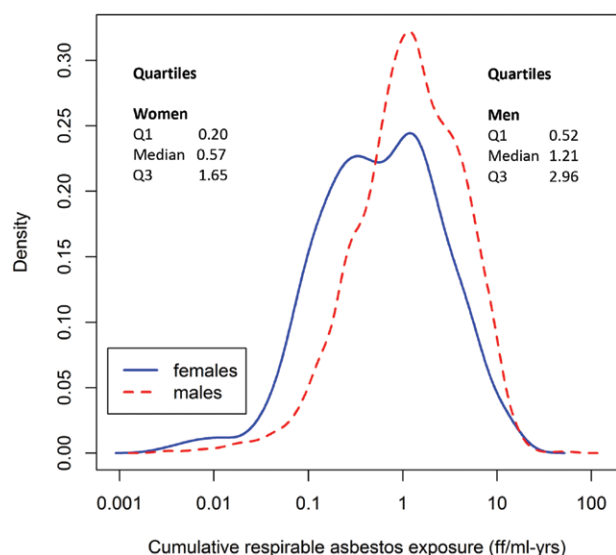
### Asbestos Exposure

At some time, 44% of cases (51% in men, 15% in women) and 35% of controls (41% in men, 11% in women) had been exposed to asbestos at their workplace. The exposure prevalence among male blue-collar workers was 63% in cases and 58% in controls. eTable 1 (<http://links.lww.com/EDE/B144>) displays the period for which asbestos exposure was assigned to workers in different studies; it does not reflect when asbestos was banned as some jobs continued to be exposed after the ban. Prevalence of asbestos exposure among control subjects by study and sex, omitting “laborers not elsewhere classified” and restricting to DOM-JEM high levels of asbestos exposure, is displayed in the eTables 2 and 3 (<http://links.lww.com/EDE/B144>).

Figure 1 shows the distribution of fiber-years in exposed control subjects by sex. Overall, women were exposed to

**TABLE 1.** Descriptive Characteristics of the Study Participants (16,901 Lung Cancer Cases, 20,965 Control Subjects) by Asbestos Exposure Status

		Asbestos Exposed				Nonexposed to Asbestos			
		Cases		Controls		Cases		Controls	
Characteristic	Exposure Category	No.	% or Mean (SD)	No.	% or Mean (SD)	No.	% or Mean (SD)	No.	% or Mean (SD)
Men									
Age	Mean (SD)	6,958	61.8 (9.0)	6,802	61.5 (9.3)	6,647	62.9 (8.9)	9,649	62.1 (9.5)
Age categorized (years)	<45	269	3.9	367	5.4	217	3.3	528	5.5
	45–64	3,728	53.6	3,514	51.7	3,257	49.0	4,668	48.4
	65+	2,961	42.6	2,921	42.9	3,173	47.7	4,453	46.1
Smoking status	Never-smoker	205	2.9	1,546	22.7	285	4.3	2,891	30.0
	Former smoker	2,422	34.8	3,135	46.1	2,365	35.6	4,193	43.5
	Current smoker	4,331	62.2	2,121	31.2	3,997	60.1	2,565	26.6
Smoking pack-years (current and former smokers)	<10	296	4.4	1,001	19.0	312	4.9	1,465	21.7
	10–19	647	9.6	1,024	19.5	568	8.9	1,399	20.7
	20+	5,810	86.0	3,231	61.5	5,482	86.2	3,894	57.6
Years-since-quitting smoking (former smokers)	2–7	905	37.4	510	16.3	841	35.6	708	16.9
	8–15	698	28.8	760	24.2	657	27.8	931	22.2
	16–25	505	20.9	884	28.2	539	22.8	1,211	28.9
	26+	314	13.0	981	31.3	328	13.9	1,343	32.0
Employed in “list A” job	Ever	1,380	19.8	1,020	15.0	349	5.3	304	3.2
Lung cancer cell type	Not available	38	0.5	-	-	41	0.6	-	-
	Adenocarcinoma	1,577	22.7	-	-	1,748	26.3	-	-
	Squamous-cell carcinoma	3,145	45.2	-	-	2,683	40.4	-	-
	Small-cell carcinoma	1,129	16.2	-	-	1,071	16.1	-	-
	Other/unspecified	1,069	15.4	-	-	1,104	16.6	-	-
Women									
Age	Mean (SD)	482	60.6 (10.5)	510	61.9 (10.0)	2,814	60.4 (10.1)	4,004	59.7 (11.5)
Age categorized (years)	<45	32	6.6	30	5.9	197	7.0	446	11.1
	45–64	260	54.0	252	49.4	1,518	53.9	1,920	48.0
	65+	190	39.4	228	44.7	1,099	39.1	1,638	40.9
Smoking status	Never-smoker	101	21.0	267	52.4	778	27.6	2,449	61.2
	Former smoker	75	15.6	126	24.7	570	20.3	766	19.1
	Current smoker	306	63.5	117	22.9	1,466	52.1	789	19.7
Smoking pack-years (current and former smokers)	<10	21	5.5	80	32.8	205	10.1	569	36.6
	10–19	64	16.8	43	17.6	328	16.1	377	24.3
	20–29	296	77.7	127	52.3	1,533	75.3	629	40.5
Years-since-quitting smoking (former smokers)	2–7	36	48.0	28	22.2	244	42.8	176	23.0
	8–15	16	21.3	28	22.2	160	28.1	179	23.4
	16–25	16	21.3	34	27.0	111	19.5	217	28.3
	26+	7	9.3	36	28.6	55	9.6	194	25.3
Employed in “list A” job	Ever	32	6.6	22	4.3	26	1.0	18	0.5
Lung cancer cell type	Not available	4	0.8	-	-	11	0.4	-	-
	Adenocarcinoma	178	36.9	-	-	1,249	44.4	-	-
	Squamous-cell carcinoma	117	24.3	-	-	558	19.8	-	-
	Small-cell carcinoma	94	19.5	-	-	436	15.5	-	-
	Other/unspecified	89	18.5	-	-	560	19.9	-	-



**FIGURE 1.** Kernel plot showing distributions of cumulative asbestos exposure in exposed control subjects among women and men in the SYNERGY project.

lower cumulative levels (median, 0.57 ff/ml-years) than men (median, 1.21 ff/ml-years).

### Lung Cancer Risk Associated with Asbestos Exposure

The ORs for lung cancer associated with ever occupational asbestos exposure changed after adjustment for smoking and for other occupational exposures. In men, OR1, from the model adjusted for age group and study, was 1.43 (95% CI, 1.37, 1.50); OR2, also adjusted for smoking, was 1.29 (95% CI, 1.22, 1.36); and OR3, also adjusted for ever-employment in a “list A” job, was 1.24 (95% CI, 1.18, 1.31). Among women, OR1 was 1.37 (95% CI, 1.19, 1.58), OR2 was 1.13 (95% CI, 0.97, 1.33), and OR3 was 1.12 (95% CI, 0.95, 1.31). In Table 2 and all subsequent analyses, we present OR3 unless otherwise stated.

In men, ORs across most exposure categories by duration and cumulative exposure were increased compared with the reference category of never-exposed to asbestos (Table 2). Only the first quartile of cumulative dose (<0.5 ff/ml-years) showed no increased risk (OR, 1.06; 95% CI, 0.96, 1.16). “Time since last exposure” was also adjusted for duration and showed ORs between 1.10 and 1.19, with no time trend ( $P = 0.44$ ).

In women, based on smaller numbers of exposed subjects and a lower median exposure level, no increased ORs were observed (Table 2).

Nonparametric exposure–response analyses showed marginal support for a nonlinear exposure–response association among men, which was larger for models with longer lag-times, while the exposure–response was linear among women (Figure 2).

Study-specific results showed no heterogeneity between studies, with  $I^2 = 0\%$  in both men and women (eFigures 4 and

5; <http://links.lww.com/EDE/B144>, which show forest plots and heterogeneity tests).

The exposure–response slope ( $K_L \times 100$ ) among all men was 6.1 (95% CI, 4.1, 8.1) and among only blue-collar workers was 3.3 (95% CI, 1.5, 5.0).

### Sensitivity Analyses

The OR for ever-exposure to asbestos in men remained stable when omitting one study at a time; the highest OR, 1.27 (95% CI, 1.19, 1.34), was observed when LUCAS was omitted, and the lowest OR, 1.21 (95% CI, 1.14, 1.29), when AUT-Munich was omitted (data not shown).

When excluding one industry or occupation at a time, the lung cancer risk in the highest quartile of cumulative exposure in men (>2.8 ff/ml-years; OR, 1.38; 95% CI, 1.27, 1.50) remained elevated, as follows: excluding asbestos manufacturing (OR, 1.44; 95% CI, 1.32, 1.57), excluding construction (OR, 1.41; 95% CI, 1.27, 1.57), excluding mining (OR, 1.36; 95% CI, 1.25, 1.49), excluding metal work (OR, 1.43; 95% CI, 1.31, 1.56), excluding transportation (OR, 1.42; 95% CI, 1.30, 1.56), or excluding vehicle mechanic (OR, 1.46; 95% CI, 1.34, 1.60) (data not shown).

Further sensitivity analyses in men (Table 3) on cumulative asbestos exposure showed that stratifying the analyses by studies with population- and hospital-based controls made a difference; ORs in studies with hospital-based controls were generally lower and more imprecise. Restricting the study population to blue-collar workers resulted in a systematic attenuation of the OR by about 10%–15%, although the significant exposure–response trend persisted. Restricting the study population to workers who started working after 1960 also lowered the ORs, whereas excluding “laborers not elsewhere classified” (ISCO 9–99.10) did not markedly influence the overall results.

### Lung Cancer Risk Associated with Cumulative Asbestos Exposure, Stratified by Major Histological Subtype and Smoking Status

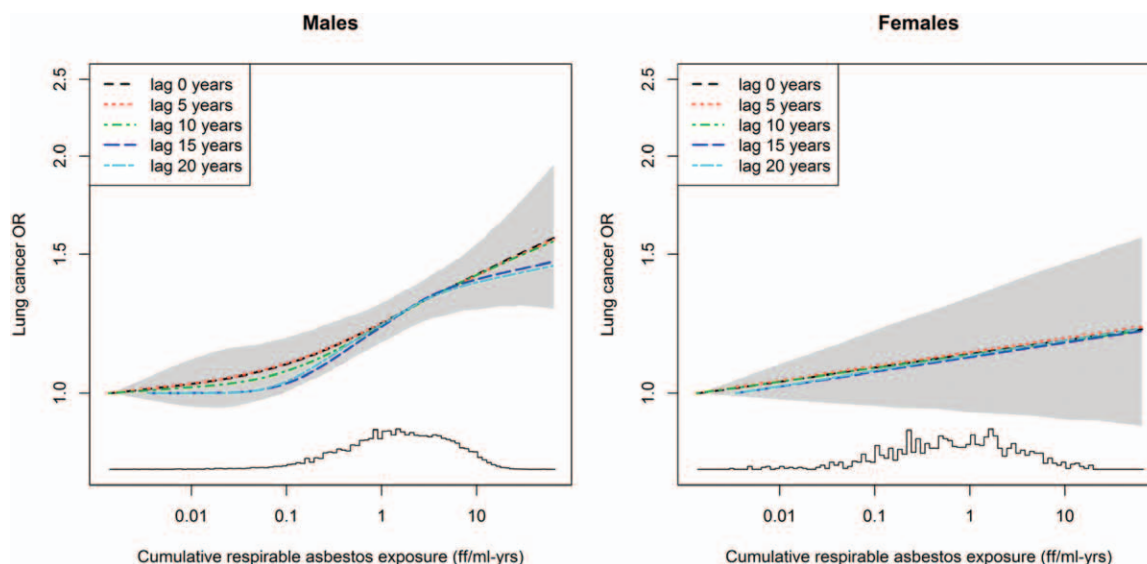
Table 4 shows ORs associated with cumulative asbestos exposure by major lung cancer subtype and by smoking status. Occupational asbestos exposure in men was associated with an increased lung cancer risk among never-smokers, former smokers, and current smokers. Never-smokers with exposure above the median (>1.2 ff/ml-years) had slightly higher ORs than former or current smokers, particularly for small-cell carcinoma (OR, 2.73; 95% CI, 1.39, 5.35). ORs were higher for squamous and small-cell carcinoma than for lung adenocarcinoma ( $P = 0.11$  for the likelihood ratio test of homogeneity from the multinomial logistic regression model when these three subtypes were included). In women, stratifying by smoking status and lung cancer subtype revealed associations in subgroups. Among current smokers, we observed associations of asbestos exposure with all lung cancer subtypes, with all ORs increased approximately two-fold. In former smokers, none of the associations was increased; and among never-smokers,

**TABLE 2.** Lung Cancer ORs and 95% CIs in Relation to Indices of Occupational Asbestos Exposure in the SYNERGY Study, 1985–2010

Indices of Occupational Asbestos Exposure	Exposure Category	Men			Women		
		Cases (%)	Controls (%)	OR3 <sup>a</sup> (95% CI)	Cases (%)	Controls (%)	OR3 <sup>a</sup> (95% CI)
Occupational asbestos exposure	Never	6,629 (48.8)	9,608 (58.5)	1.00 (reference)	2,717 (84.9)	3,898 (88.4)	1.00 (reference)
	Ever	6,958 (51.2)	6,802 (41.5)	1.24 (1.18, 1.31)	482 (15.1)	510 (11.6)	1.12 (0.95, 1.31)
Duration (years)	1–9	2,425 (17.8)	2,603 (15.9)	1.16 (1.07, 1.25)	303 (9.5)	300 (6.8)	1.16 (0.96, 1.42)
	10–19	1,375 (10.1)	1,374 (8.4)	1.19 (1.08, 1.30)	102 (3.2)	122 (2.8)	1.08 (0.79, 1.47)
	20–29	1,141 (8.4)	1,012 (6.2)	1.34 (1.21, 1.49)	48 (1.5)	57 (1.3)	0.94 (0.61, 1.46)
	30+	2,017 (14.8)	1,813 (11.0)	1.35 (1.25, 1.47)	29 (0.9)	31 (0.7)	1.18 (0.66, 2.11)
Test for trend, <i>P</i> value				<0.01			0.48
Excl. never asbestos exposed				<0.01			0.98
Cumulative exposure (ff/ml-years)	<0.5	1,206 (8.9)	1,593 (9.7)	1.06 (0.96, 1.16)	194 (6.1)	230 (5.2)	1.11 0.87, 1.42
	<1.2	1,624 (12.0)	1,713 (10.4)	1.26 (1.15, 1.37)	104 (3.3)	104 (2.4)	0.95 (0.69, 1.31)
	<2.8	1,840 (13.5)	1,724 (10.5)	1.25 (1.15, 1.36)	110 (3.4)	106 (2.4)	1.23 (0.90, 1.68)
	>2.8	2,288 (16.8)	1,772 (10.8)	1.38 (1.27, 1.50)	74 (2.3)	70 (1.6)	1.22 (0.84, 1.78)
Test for trend, <i>P</i> value				<0.01			0.17
Excl. never asbestos exposed				<0.01			0.82
Time since last exposure (years) <sup>b</sup>	1–4	1,646 (12.1)	1,589 (9.7)	1.17 (0.99, 1.37)	75 (2.3)	67 (1.5)	1.53 (0.93, 2.53)
	5–9	838 (6.2)	765 (4.7)	1.18 (1.00, 1.40)	34 (1.1)	36 (0.8)	1.33 (0.72, 2.47)
	10–19	1,297 (9.5)	1,171 (7.1)	1.19 (1.03, 1.37)	79 (2.5)	81 (1.8)	1.27 (0.81, 2.00)
	20–29	924 (6.8)	917 (5.6)	1.12 (0.98, 1.27)	85 (2.7)	106 (2.4)	1.10 (0.75, 1.61)
	30–39	1,127 (8.3)	1,162 (7.1)	1.16 (1.04, 1.29)	85 (2.7)	106 (2.4)	0.94 (0.65, 1.34)
	40+	1,126 (8.3)	1,198 (7.3)	1.10 (0.99, 1.22)	124 (3.9)	114 (2.6)	1.32 (0.96, 1.80)
Test for trend, <i>P</i> value				0.44			0.44
Excl. never asbestos exposed				0.44			0.44

<sup>a</sup>OR3 is adjusted for study, age group, smoking (pack-years, time-since-quitting smoking), and list A jobs.

<sup>b</sup>OR3 in “time since last asbestos exposure” is in addition adjusted for duration (continuous) of asbestos exposure.



**FIGURE 2.** Exposure–response relationship among men and women for cumulative asbestos exposure with different lag periods applied and 95% CIs around the 0-year-lag curve, adjusted for study, age group, cigarette pack-years, time-since-quitting smoking, and ever-employment in a “list A” job. The histograms on the x axis show how the study populations are distributed.



**TABLE 3.** Lung Cancer ORs and 95% CIs in Relation to Cumulative Asbestos Exposure in Restricted Strata (Sensitivity Analyses) of Men in the SYNERGY Study, 1985–2010

Cumulative Asbestos Exposure (ff/ml-years)	Studies with Population Controls		Studies with Hospital Controls		Restricting the Study Base to Blue-collar Workers	
	ca/co	OR3 <sup>a</sup> (95% CI)	ca/co	OR3 <sup>a</sup> (95% CI)	ca/co	OR3 <sup>a</sup> (95% CI)
Unexposed	5,136/7,894	1	1,609/1,714	1	4,067/4,721	1
<0.5	1,048/1,447	1.05 (0.95, 1.17)	188/146	1.14 (0.89, 1.47)	1,183/1,552	0.93 (0.84, 1.03)
<1.2	1,376/1,502	1.28 (1.16, 1.40)	267/211	1.26 (1.02, 1.57)	1,588/1,653	1.12 (1.02, 1.23)
<2.8	1,428/1,328	1.31 (1.19, 1.44)	429/396	1.06 (0.90, 1.26)	1,810/1,663	1.11 (1.02, 1.22)
>2.8	1,563/1,161	1.50 (1.35, 1.66)	611/725	1.16 (1.00, 1.34)	2,249/1,724	1.23 (1.12, 1.35)
Test for trend, <i>P</i> value		<0.01		0.02		<0.01
Excl. never asbestos exposed		<0.01		0.82		<0.01

Cumulative Asbestos Exposure (ff/ml-years)	Restricted to Workers Started 1960 or Later		Excluding Laborers Not Elsewhere Classified (ISCO 9–99.10)	
	ca/co	OR3 <sup>a</sup> (95% CI)	ca/co	OR3 <sup>a</sup> (95% CI)
Unexposed	1,709/2,836	1	6,572/9,544	1
<0.5	395/570	0.96 (0.80, 1.14)	1,032/1,359	1.06 (0.96, 1.17)
<1.2	411/463	1.19 (1.00, 1.42)	1,318/1,443	1.23 (1.12, 1.35)
<2.8	405/395	1.40 (1.17, 1.67)	1,461/1,372	1.28 (1.16, 1.40)
>2.8	279/267	1.27 (1.02, 1.59)	1,752/1,442	1.33 (1.21, 1.46)
Test for trend, <i>P</i> value		<0.01		<0.01
Excl. never asbestos exposed		0.06		<0.01

<sup>a</sup>OR3 is adjusted for study, age group, smoking (pack-years, time-since-quitting smoking), and list A jobs. ca indicates cases; co, controls.

our results showed no association for lung adenocarcinoma or squamous-cell lung cancer but a relatively strong association for small-cell lung cancer even at low levels of asbestos exposure (<1.2 ff/ml-years: OR, 3.51; 95% CI, 1.29, 9.55).

### Joint Effects of Asbestos and Smoking

Table 5 shows the joint effects of asbestos exposure and smoking, overall and by lung cancer subtype. In men, the joint effect of smoking and asbestos was more than additive for all lung cancer subtypes, with a higher excess risk due to interaction for squamous- and small-cell lung carcinoma than for lung adenocarcinoma, whereas there was no deviation from a multiplicative scale ( $P = 0.10$ – $0.90$ ). Patterns were similar in women, but the RERIs were not significantly different from 0, except for squamous-cell lung carcinoma. The strong association between asbestos exposure and small-cell lung carcinoma in never-smokers resulted in a submultiplicative interaction with smoking, in women ( $P = 0.01$ ) but not in men ( $P = 0.10$ ). A complementary table including ORs for models with and without interaction between occupational asbestos exposure and smoking is shown in eTable 6 (<http://links.lww.com/EDE/B144>).

### DISCUSSION

We investigated the quantitative association between occupational asbestos exposure and lung cancer risk in the SYNERGY project by sex, smoking status, and lung cancer

subtype. Increasing duration and increasing cumulative asbestos exposure were associated with an increasing lung cancer risk in men. Moreover, the increased lung cancer risk in men was observed in never-smokers, former smokers, and current smokers, and for all three major lung cancer subtypes. Sensitivity analyses revealed that results were not driven by exposure in any particular industry or study. Women were exposed to lower levels of asbestos than men (median, 0.57 vs. 1.21 ff/ml-years), which may explain the weaker association with lung cancer among women. The interaction between asbestos exposure and smoking was more than additive for all major lung cancer subtypes among men and for squamous-cell lung carcinoma among women; moreover, the interaction between asbestos and smoking among men did not deviate from multiplicativity. The results from our pooled analysis of case-control studies are in broad agreement with those obtained by Wraith and Mengersen<sup>12</sup> in a meta-analysis of industry-based cohort and case-control studies, although our study adds results on the interaction between smoking and asbestos by lung cancer subtype.

Lagged exposure estimates generated very similar results (not shown) to those of unlagged estimates; a possible explanation is that the relative exposure distribution remained the same because most exposed subjects were exposed to no or low exposure levels in recent decades, particularly after the implementation of asbestos bans in the different countries.

**TABLE 4.** Lung Cancer ORs and 95% CIs in Relation to Cumulative Asbestos Exposure Stratified by Lung Cancer Cell Type and Smoking Status Among Men and Women in the SYNERGY Study, 1985–2010

Lung Cancer Cell Type	Asbestos Exp. (ff/ml-years)	Never-smokers <sup>a</sup>			Former Smokers <sup>b</sup>			Current Smokers <sup>c</sup>		
		Control	Case	OR (95% CI)	Control	Case	OR (95% CI)	Control	Case	OR (95% CI)
Men										
All lung cancer	Unexposed	2,875	283	1	4,185	2,363	1	2,548	3,983	1
	<1.2	801	103	1.31 (1.02, 1.69)	1,526	1,001	1.20 (1.08, 1.33)	979	1,726	1.12 (1.01, 1.24)
	>1.2	745	102	1.51 (1.16, 1.97)	1,609	1,421	1.38 (1.25, 1.52)	1,142	2,605	1.21 (1.10, 1.34)
Test for trend, <i>P</i> value				<0.01	<0.01			<0.01		
Excl. never asbestos exposed				0.04	<0.01			0.06		
Adenocarcinoma	Unexposed	2,875	103	1	4,185	663	1	2,548	976	1
	<1.2	801	34	1.13 (0.75, 1.72)	1,526	268	1.10 (0.93, 1.30)	979	425	1.05 (0.90, 1.22)
	>1.2	745	40	1.52 (1.00, 2.29)	1,609	340	1.30 (1.11, 1.53)	1,142	470	1.06 (0.92, 1.23)
Test for trend, <i>P</i> value				0.07	<0.01			0.23		
Excl. never asbestos exposed				<0.01	<0.01			<0.01		
Squamous cell	Unexposed	2,875	78	1	4,185	984	1	2,548	1,617	1
	<1.2	801	36	1.72 (1.13, 2.61)	1,526	441	1.32 (1.14, 1.52)	979	707	1.18 (1.04, 1.34)
	>1.2	745	27	1.27 (0.78, 2.06)	1,609	670	1.47 (1.29, 1.67)	1,142	1,264	1.29 (1.16, 1.45)
Test for trend, <i>P</i> value				0.04	<0.01			<0.01		
Excl. never asbestos exposed				<0.01	<0.01			<0.01		
Small cell	Unexposed	2,875	26	1	4,185	312	1	2,548	729	1
	<1.2	801	12	1.55 (0.76, 3.18)	1,526	123	1.12 (0.89, 1.42)	979	297	1.05 (0.89, 1.24)
	>1.2	745	18	2.73 (1.39, 5.35)	1,609	192	1.36 (1.10, 1.68)	1,142	487	1.22 (1.05, 1.43)
Test for trend, <i>P</i> value				<0.01	<0.01			0.02		
Excl. never asbestos exposed				<0.01	<0.01			<0.01		
Women										
All lung cancer	Unexposed	2,377	755	1	749	557	1	772	1,405	1
	<1.2	139	42	1.06 (0.72, 1.54)	101	49	0.71 (0.46, 1.08)	94	207	1.22 (0.90, 1.66)
	>1.2	128	59	0.95 (0.67, 1.34)	25	26	1.09 (0.57, 2.07)	23	99	2.05 (1.24, 3.40)
Test for trend, <i>P</i> value				0.95	0.18			<0.01		
Excl. never asbestos exposed				0.96	0.32			0.16		
Adenocarcinoma	Unexposed	2,377	458	1	749	249	1	772	502	1
	<1.2	139	22	0.92 (0.57, 1.50)	101	15	0.43 (0.23, 0.79)	94	70	1.11 (0.76, 1.63)
	>1.2	128	29	0.73 (0.47, 1.15)	25	13	1.64 (0.74, 3.66)	23	29	2.08 (1.12, 3.88)
Test for trend, <i>P</i> value				0.19	0.14			0.08		
Excl. never asbestos exposed				<0.01	<0.01			<0.01		
Squamous cell	Unexposed	2,377	99	1	749	129	1	772	306	1
	<1.2	139	5	0.92 (0.36, 2.38)	101	18	1.04 (0.53, 2.04)	94	49	1.33 (0.86, 2.07)
	>1.2	128	8	0.77 (0.34, 1.73)	25	6	0.77 (0.26, 2.26)	23	31	2.20 (1.15, 4.21)
Test for trend, <i>P</i> value				0.56	0.63			<0.01		
Excl. never asbestos exposed				<0.01	<0.01			<0.01		
Small cell	Unexposed	2,377	39	1	749	72	1	772	313	1
	<1.2	139	5	3.51 (1.29, 9.55)	101	5	0.53 (0.18, 1.55)	94	42	1.15 (0.73, 1.83)
	>1.2	128	11	3.71 (1.75, 7.86)	25	2	0.30 (0.04, 2.07)	23	29	2.15 (1.12, 4.12)
Test for trend, <i>P</i> value				<0.01	0.05			0.03		
Excl. never asbestos exposed				<0.01	<0.01			<0.01		

<sup>a</sup>OR in never-smokers adjusted for study, age group, and list A jobs.<sup>b</sup>OR in former smokers adjusted for study, age group, list A jobs, cigarette pack-years, and time-since-quitting smoking.<sup>c</sup>OR in current smokers adjusted for study, age group, list A jobs, and cigarette pack-years.

Also, many workers had been retired for many years when they were diagnosed with lung cancer, and therefore their exposure did not change much, even when lagged.

The quality of the exposure assessment has a strong influence on the estimation of the exposure–response association.<sup>23</sup> So far, quantitative estimates based on measurements

**TABLE 5.** Lung Cancer ORs and 95% CIs, *P* Value for Multiplicative Interaction and RERI and 95% CI in Relation to Occupational Asbestos Exposure and Smoking Among Men and Women Overall and by Major Lung Cancer Cell Types, in the SYNERGY Study, 1985–2010

Exposure Status	All Lung Cancer Types			Lung Adenocarcinoma		Squamous-cell Carcinoma		Small-cell Carcinoma	
	Controls	Cases	OR <sup>a</sup> (95% CI) <sup>b</sup>	Cases	OR <sup>a</sup> (95% CI) <sup>b</sup>	Cases	OR <sup>a</sup> (95% CI) <sup>b</sup>	Cases	OR <sup>a</sup> (95% CI) <sup>b</sup>
<b>Men</b>									
Never-smoker and never asbestos	2,875	283	1 (reference)	103	1 (reference)	78	1 (reference)	26	1 (reference)
Never-smoker and asbestos	1,546	205	1.26 (1.04, 1.53)	74	1.29 (0.95, 1.75)	63	1.35 (0.96, 1.90)	30	1.97 (1.16, 3.35)
Ever smoker and never asbestos	6,733	6,346	9.23 (8.13, 10.5)	1,639	6.71 (5.47, 8.24)	2,601	13.2 (10.5, 16.7)	1,041	16.8 (11.3, 24.9)
Ever smoker and asbestos	5,256	6,753	11.9 (10.5, 13.6)	1,503	7.91 (6.42, 9.74)	3,082	18.3 (14.5, 23.0)	1,099	21.1 (14.3, 31.3)
<i>P</i> value multiplicative interaction			0.82		0.59		0.90		0.10
RERI <sup>c</sup> with linear model			2.44 (1.89, 3.08)		0.92 (0.16, 1.59)		4.75 (3.57, 6.55)		3.18 (1.29, 5.92)
<b>Women</b>									
Never-smoker and never asbestos	2,377	755	1 (reference)	458	1 (reference)	99	1 (reference)	39	1 (reference)
Never-smoker and asbestos	267	101	1.00 (0.78, 1.29)	51	0.84 (0.60, 1.16)	13	0.92 (0.50, 1.70)	16	3.40 (1.85, 6.26)
Ever smoker and never asbestos	1,521	1,962	4.57 (4.08, 5.12)	751	2.72 (2.36, 3.13)	435	8.90 (7.00, 11.3)	385	17.1 (12.1, 24.1)
Ever smoker and asbestos	243	381	6.26 (5.14, 7.62)	127	3.31 (2.57, 4.28)	104	13.3 (9.49, 18.5)	78	24.3 (15.8, 37.4)
<i>P</i> value multiplicative interaction			0.05		0.07		0.16		0.01
RERI <sup>c</sup> with linear model			1.69 (0.64, 3.11)		0.76 (−0.11, 1.68)		4.44 (1.02, 9.16)		4.85 (−2.23, 14.21)

<sup>a</sup>OR adjusted for study, age group, and “list A” jobs.<sup>b</sup>CI is based on 1,000 bootstrap samples.<sup>c</sup>RERI indicates relative excess risk due to interaction.

have been obtained mainly from industrial cohort studies.<sup>24</sup> We assessed asbestos exposure applying SYN-JEM for general population studies, a newly created job-exposure matrix based on quantitative workplace measurements from Europe and Canada.

Strengths of this study include the large study population, a large proportion of face-to-face interviews conducted by trained interviewers, a large proportion of control subjects recruited from the general population, a comprehensive adjustment for smoking, and an innovative and objective method, supported by actual measurements, for assessing asbestos exposure quantitatively in general population studies.

Limitations of the study are that asbestos fiber type and dimensions could not be taken into account, because almost all measurements (>95%) were determined by phase-contrast microscopy, which does not allow the fiber type to be distinguished or to identify fibers with width <0.25  $\mu$ m. For lung cancer, scientific uncertainty remains on how much risks differ in magnitude by fiber type, but recent evidence suggests it is less than previously assumed.<sup>24,25</sup>

Misclassification of exposure is assumed to have occurred, but it can also be assumed to be nondifferential provided reporting of job titles did not differ systematically in cases and controls. Differential reporting of job titles is not likely; therefore, the use of a job-exposure matrix for exposure assessment is unlikely to have created spurious associations.

Further limitations include that measurements were not done for individual study subjects, resulting in assignments of

average exposure levels to job titles and not to individuals, and leading to assignment of the same exposure level to all workers sharing the same job code in a particular year, irrespective of exposure variability between workers in the same job. This results in a Berkson-type error, which usually does not bias the point estimate but increases the variance and therefore leads to reduced precision.<sup>26</sup>

Some jobs or unspecific job codes may substantially influence the prevalence of exposure if used extensively in a study. For example, many jobs were coded as “laborers not elsewhere classified” (ISCO 9–99.10); in SYN-JEM, this job was assigned low exposure to asbestos. The use of the ISCO 9–99.10 job code does not necessarily reflect poor quality of interviews or coding, as it could also signify a high prevalence of low specialization of laborers and/or low technology in certain industries. Indeed, the asbestos prevalence decreased by >10% in some studies when excluding “laborers not elsewhere classified.” Nevertheless, excluding “laborers not elsewhere classified” from the risk analyses did not change the overall results, possibly because 9–99.10 jobs represented only 1.8% of the total working time.

Another limitation is that half of the 36,000 personal measurements collected were available for the production of asbestos cement and asbestos textiles, job titles that were rare in our study population or too specific to be captured by the ISCO code.<sup>15</sup> In the SYNERGY study population, only 46 subjects (32 cases, 14 controls) had ever worked as asbestos cement product makers. The number of available data points for the remaining, more prevalent jobs was more limited.

Exposure assessment according to SYN-JEM resulted in high prevalence of ever occupational asbestos exposure compared with the original studies.<sup>27–30</sup> However, the prevalence of asbestos exposure decreased substantially, from 41.3% to 6.4% among male controls and from 11.3% to 0.3% among female controls, when only the DOM-JEM high-exposure jobs were considered, which confirms that the vast majority of exposed workers were employed in jobs with low average exposure levels.

About 60% of all blue-collar workers were rated as ever-exposed to asbestos. This implies that they may have been exposed to many other agents at their workplaces. Although we controlled for occupations with known exposure to pulmonary carcinogens, there may still remain some uncertainty about residual confounding by other occupational hazards. Notably, the association between asbestos exposure and lung cancer risk was weaker when restricting to blue-collar workers. This may be due to a reduction in exposure contrast. Alternatively, the background lung cancer risk may be higher in blue-collar workers than in white-collar workers due to other agents in the workplace and various socioeconomic factors.<sup>31–33</sup> Also, selection bias has to be taken into account due to a lower participation rate of blue-collar workers among population controls.<sup>34</sup> The hospital-based studies showed lower ORs for asbestos exposure than the population-based studies, which may be explained by a combination of factors, including choice of control diseases, geographical location (possibly reflecting different exposure patterns), study size, or other factors.

Levels of occupational asbestos exposure in this pooled analysis of general population studies were lower (range, 0.0023–64.6 ff/ml-years in male controls) than in the 18 industrial cohort studies (range, 0.11–4,710 ff/ml-years) included in a recent review of exposure–response relationships.<sup>24</sup> A probable reason for the rather low levels of occupational asbestos exposure seen in SYNERGY is that major “asbestos occupations” such as asbestos cement product makers and asbestos textile workers are rare in a general population study setting. Instead, exposures in SYNERGY represent a wider exposure range, with very few workers exposed to high levels and most being downstream users or indirectly exposed workers occasionally exposed or exposed to lower concentrations of fibers only.

Our large dataset may be particularly informative to explore the shape of the exposure–response function in the low-dose range. An additional advantage was stratification or detailed adjustment for smoking. In our analysis of pooled general population studies, the exposure–response slope estimated as excess risk per 100 fiber-years ( $K_L * 100$ ) was 6.1 (95% CI, 4.1, 8.1) in men overall and 3.3 (95% CI, 1.5, 5.0) among male blue-collar workers. The  $K_L * 100$  slope was flat in a recent meta-analysis of 18 occupational cohort studies ( $K_L * 100$ , 0.13; 95% CI, 0.04, 0.22), while it was rather steep ( $K_L * 100$ , 15.5; 95% CI, 1.13, 29.87) in LUCAS, the general

population study from Stockholm, which also is part of SYN-ERGY.<sup>24,35</sup> The SYNERGY  $K_L * 100$  estimate in blue-collar workers (3.3; 95% CI, 1.5, 5.0) is still considerably higher than the estimate from the industrial cohorts in Lenters’ paper, which were considered to have good-quality exposure assessment (e.g., seven cohort studies with >30% coverage of exposure data;  $K_L * 100$ , 0.27; 95% CI, 0.08, 0.46). A possible reason for the steeper slope we observed in general population studies is that we could assess the full occupational history, resulting in a more distinct exposure contrast, and that we could identify a substantial proportion of truly nonexposed. However, there is little evidence in the literature regarding the shape of the exposure–response curve at low levels of exposure.<sup>25,36</sup> Our dataset is less informative regarding the high-dose range; only 90 of 29,997 male participants were assessed as exposed to  $\geq 15$  fiber-years. The exposure–response results presented here are based on exposure–response modeling agreed upon a priori. This model assumes  $\ln(\text{OR})$  is proportional to  $\ln(\text{exposure})$ . Other exposure–response models will result in different risk estimates.

We observed a stronger association between asbestos exposure and small-cell lung carcinoma among never-smokers in both men and women. This is noteworthy because small- and squamous-cell lung carcinomas occur almost exclusively in cigarette smokers; in SYNERGY, only 4% of small- and squamous-cell cases were never-smokers, whereas 14% of lung adenocarcinoma cases were never-smokers. However, we cannot rule out biased recall of smoking habits.<sup>37</sup>

Some misclassification of the histological subtypes of lung cancer is likely; one of the studies (HdA) included in SYNERGY assessed diagnostic agreement between pathologists and found a kappa of 0.54 (95% CI, 0.49, 0.58).<sup>14</sup> Small-cell lung cancer was best classified, followed by squamous-cell lung cancer and lung adenocarcinoma. Most misclassification was between squamous-cell lung cancer and lung adenocarcinoma. Thus, our results for the major lung cancer subtypes should be interpreted with caution.

Our results show an excess risk of lung cancer and its subtypes at relatively low levels of cumulative exposure ( $>0.5$  ff/ml-years), which persisted at least up to 40 years after last exposure. Furthermore, the slope of the exposure–response relationship seemed steeper in this exposure range than at higher (and previously studied) levels. Together, this implies that the future burden of disease due to asbestos exposure may be underestimated.

## ACKNOWLEDGMENTS

*The authors wish to thank Guillermo Fernandez-Tardon, hygienist at the University Institute of Oncology of Asturias - Cajastur Social Programme (IUOPA), in Oviedo, Asturias.*



## REFERENCES

- Albin M, Magnani C, Krstev S, Rapiti E, Shefer I. Asbestos and cancer: an overview of current trends in Europe. *Environ Health Perspect.* 1999;107(suppl 2):289–298.
- Committee on asbestos: selected health effects. *Asbestos: Selected Cancers.* Washington D.C.: The National Academies Press; 2006.
- IARC working group on the evaluation of carcinogenic risks to humans, which met in Lyon 17–24 March 2009. A review of human carcinogens. Vol. 100 C. Metals, Arsenic, Dusts and Fibres. 2011. World Health Organization; International Agency for Research on Cancer. IARC Monographs on the Evaluation of Carcinogenic Risks to Humans.
- WHO. Elimination of asbestos related diseases. WHO/SDE/OEH/06.03. 2006. Geneva, World Health Organization.
- Ferlay J, Soerjomataram I, Ervik M, et al. GLOBOCAN 2012 v1.2, cancer incidence and mortality worldwide: IARC CancerBase No. 11 [Internet]. Lyon, France: International Agency for Research on Cancer; 2015. Available at: <http://globocan.iarc.fr>. Accessed January 6, 2017.
- Parkin DM, Tyczynski JE, Boffetta P, Samet JM, Shields P, Caporaso NE. Lung cancer epidemiology and etiology. In: Travis WD, Brambilla E, Müller-Hermelink HK, Harris CC, eds. *Pathology and Genetics of Tumours of the Lung, Pleura, Thymus and Heart.* Lyon: IARC Press; 2004:12–15.
- Straif K. The burden of occupational cancer. *Occup Environ Med.* 2008;65:787–788.
- Selikoff IJ, Hammond EC, Churg J. Asbestos exposure, smoking, and neoplasia. *JAMA.* 1968;204:106–112.
- Erren TC, Jacobsen M, Piekarski C. Synergy between asbestos and smoking on lung cancer risks. *Epidemiology.* 1999;10:405–411.
- IARC Working Group on the Evaluation of Carcinogenic Risks to Humans, which met in Lyon 11–18 June 2004. A review of human carcinogens. Vol. 83. Tobacco smoke and involuntary smoking. 2004. WHO, International Agency for Research on Cancer IARC working group. Lyon, France, IARC. IARC Monographs on the evaluation of carcinogenic risks to humans.
- Lee PN. Relation between exposure to asbestos and smoking jointly and the risk of lung cancer. *Occup Environ Med.* 2001;58:145–153.
- Wraith D, Mengersen K. Assessing the combined effect of asbestos exposure and smoking on lung cancer: a Bayesian approach. *Stat Med.* 2007;26:1150–1169.
- Liddell FD. The interaction of asbestos and smoking in lung cancer. *Ann Occup Hyg.* 2001;45:341–356.
- Stang A, Pohlmann H, Müller KM, Jahn I, Giersiepen K, Jöckel KH. Diagnostic agreement in the histopathological evaluation of lung cancer tissue in a population-based case-control study. *Lung Cancer.* 2006;52:29–36.
- Peters S, Vermeulen R, Olsson A, et al. Development of an exposure measurement database on five lung carcinogens (ExpoSYN) for quantitative retrospective occupational exposure assessment. *Ann Occup Hyg.* 2012;56:70–79.
- Peters S, Vermeulen R, Portengen L, et al. Modelling of occupational respirable crystalline silica exposure for quantitative exposure assessment in community-based case-control studies. *J Environ Monit.* 2011;13:3262–3268.
- Peters S, Kromhout H, Portengen L, et al. Sensitivity analyses of exposure estimates from a quantitative job-exposure matrix (SYN-JEM) for use in community-based studies. *Ann Occup Hyg.* 2013;57:98–106.
- Peters S, Vermeulen R, Portengen L, et al. SYN-JEM: a quantitative job-exposure matrix for five lung carcinogens. *Ann Occup Hyg.* 2016;60:795–811.
- Ahrens W, Merletti F. A standard tool for the analysis of occupational lung cancer in epidemiologic studies. *Int J Occup Environ Health.* 1998;4:236–240.
- Mirabelli D, Chiusolo M, Calisti R, et al. Database of occupations and industrial activities that involve the risk of pulmonary tumors. *Epidemiol Prev.* 2001;25:215–221.
- Higgins JP, Thompson SG, Deeks JJ, Altman DG. Measuring inconsistency in meta-analyses. *BMJ.* 2003;327:557–560.
- Richardson DB, Kaufman JS. Estimation of the relative excess risk due to interaction and associated confidence bounds. *Am J Epidemiol.* 2009;169:756–760.
- Armstrong BG. Effect of measurement error on epidemiological studies of environmental and occupational exposures. *Occup Environ Med.* 1998;55:651–656.
- Lenters V, Vermeulen R, Dogger S, et al. A meta-analysis of asbestos and lung cancer: is better quality exposure assessment associated with steeper slopes of the exposure-response relationships? *Environ Health Perspect.* 2011;119:1547–1555.
- van der Bij S, Koffijberg H, Lenters V, et al. Lung cancer risk at low cumulative asbestos exposure: meta-regression of the exposure-response relationship. *Cancer Causes Control.* 2013;24:1–12.
- Kreienbrock L. Environmental epidemiology, chapt. III.3. In: Ahrens W, Pigeot I, eds. *Handbook of Epidemiology.* Berlin Heidelberg, Germany: Springer; 2007.
- Bröske-Hohlfeld I, Möhner M, Pohlmann H, et al. Occupational lung cancer risk for men in Germany: results from a pooled case-control study. *Am J Epidemiol.* 2000;151:384–395.
- Carel R, Olsson AC, Zaridze D, et al. Occupational exposure to asbestos and man-made vitreous fibres and risk of lung cancer: a multicentre case-control study in Europe. *Occup Environ Med.* 2007;64:502–508.
- Gustavsson P, Jakobsson R, Nyberg F, Pershagen G, Järup L, Schéele P. Occupational exposure and lung cancer risk: a population-based case-referent study in Sweden. *Am J Epidemiol.* 2000;152:32–40.
- Pintos J, Parent ME, Rousseau MC, Case BW, Siemiatycki J. Occupational exposure to asbestos and man-made vitreous fibers, and risk of lung cancer: evidence from two case-control studies in Montreal, Canada. *J Occup Environ Med.* 2008;50:1273–1281.
- Sidorchuk A, Agardh EE, Aremu O, Hallqvist J, Allebeck P, Moradi T. Socioeconomic differences in lung cancer incidence: a systematic review and meta-analysis. *Cancer Causes Control.* 2009;20:459–471.
- Hrubá F, Fabiánová E, Bencko V, et al. Socioeconomic indicators and risk of lung cancer in Central and Eastern Europe. *Cent Eur J Public Health.* 2009;17:115–121.
- van Loon AJ, Goldbohm RA, van den Brandt PA. Lung cancer: is there an association with socioeconomic status in The Netherlands? *J Epidemiol Community Health.* 1995;49:65–69.
- Richiardi L, Boffetta P, Merletti F. Analysis of nonresponse bias in a population-based case-control study on lung cancer. *J Clin Epidemiol.* 2002;55:1033–1040.
- Gustavsson P, Nyberg F, Pershagen G, Schéele P, Jakobsson R, Plato N. Low-dose exposure to asbestos and lung cancer: dose-response relations and interaction with smoking in a population-based case-referent study in Stockholm, Sweden. *Am J Epidemiol.* 2002;155:1016–1022.
- Nielsen LS, Bælum J, Rasmussen J, et al. Occupational asbestos exposure and lung cancer—a systematic review of the literature. *Arch Environ Health.* 2014;69:191–206.
- Pesch B, Kendzia B, Gustavsson P, et al. Cigarette smoking and lung cancer—relative risk estimates for the major histological types from a pooled analysis of case-control studies. *Int J Cancer.* 2012;131:1210–1219.